CVER VIRTUAL COFFEE SOCIAL, 7 MARCH 2002 Henrik Stryhn

Statistical Modelling of Reviewer Scores for Abstracts

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INTRODUCTION — THE PROBLEM

Task at hand:

- based on reviewer scores, rank abstracts from highest to lowest,
- make decisions about "acceptance" of abstracts at suitable cut-off(s).

Data at hand (first round of abstract submissions for ISVEE 16):

- 119 abstracts, each scored twice (0 100 scale) by two of 27 reviewers,
- \circ reviewers assessed 1 15 abstracts (average 238/27 = 8.8).

Approaches considered to base ranking on:

- (1) simple: average score for two reviewers per abstract,
- (2) model-based: estimate abstract levels from statistical model,
- (1x) expanded simple: request extra review for selected (15) abstracts, and then use (1) with simple average across all reviewers per abstract.

Aim of this exploration: determine the feasibility of (2) and compare its results with (1) and (1x).

DATA ILLUSTRATION I

Possible data layout for 10 abstracts and 5 reviewers:

	Abstract									
Reviewer	1	2	3	4	5	6	7	8	9	10
1	\checkmark	-		-	-	-		-	-	\checkmark
2	-		-	-	-		\checkmark	-	\checkmark	-
3	\checkmark	\checkmark	-	-	\checkmark	-	-	\checkmark	-	-
4	-	-	-	\checkmark	\checkmark	\checkmark	-	-	-	\checkmark
5	-	-			-	-	-		\checkmark	-

an incomplete two-way classification

- unrealistically "nice" design (actually a "balanced" incomplete block design) with

- * equal number of reviews per reviewer (4),
- * each pair of reviewers share exactly one abstract,

which gives nice (equal precision) comparisons between reviewers in a model accounting for both abstracts and reviewers. But even in this nice layout,

- simple means and adjusted means for reviewers are not the same,
- similarly nice properties do not hold for abstracts (too little replication).

Take-away message: we cannot compensate for the incompleteness by a clever design, and dependence on the other classification variable is unavoidable.

STATISTICAL MODELS

The data layout invites a two-way ANOVA,

$$y_{ij} = \mu + lpha_i + eta_j + arepsilon_{ij}, \quad ext{where}$$

- reviewer effects (α_i) and abstract effects (β_j) are taken as either fixed or random (drawn from N(0, σ^2) distribution(s)),
- assuming all $\alpha_i = 0$ corresponds to a one-way ANOVA, method (1).

Fixed or random effects? in two-way ANOVA:

- (random): assumes effects drawn from a population, avoids estimation of a large number of individual parameters, exerts smoothing on estimation, balances abstract and reviewer effects against their respective distribution assumptions,
- (fixed): no assumptions on effects estimated freely to achieve best fit, requires estimation of a very large number of parameters with potential (near-)collinearity between them (next slide),
- (mixed): fixed effects for raters (reviewers) is common in item response models (Skrondal & Rabe-Hesketh, 2004), and may be preferable for a small number of raters or with reviewer effects not approximated well by $N(0, \sigma^2)$.

Initial focus: method (1) vs. random effects model vs. mixed model.

Data Illustration II \sim Problems

Modified data layout with extra abstracts and reviewers, in two scenarios (A) and (B):

	Abstract										
Reviewer	1	2	3	4	5	6	7	8	9	10	11
1	\checkmark	-	\checkmark	-	-	-	\checkmark	-	-		-
2	-	\checkmark	-	-	-	\checkmark	\checkmark	-	\checkmark	-	-
3			-	-	\checkmark	-	-	\checkmark	-	-	-
4	-	-	-	\checkmark	\checkmark	\checkmark	-	-	-	\checkmark	-
5	-	-	\checkmark		-	-	-	\checkmark	\checkmark	-	-
6	-	-	-	-	-	-	-	-	-	-	\checkmark
7	-	-	-	-	-	-	-	-	-	-	\checkmark

— a (fixed effects) collinearity between added reviewer and abstract effects,

- * effects of reviewers 6-7 and abstract 11 cannot be separated from each other,
- * effectively, 3 added parameters but only 2 extra observations (and essentially the same problem would occur with reviewer 6 only),
- * these types of collinearities can typically be detected from data summaries.

Note: estimation would still be possible in a random effects model!

Data Illustration II \sim Problems

Modified data layout with extra abstracts and reviewers, in two scenarios (A) and (B):

						A	bstr	act					
Reviewer	1	2	3	4	5	6	7	8	9	10	11	12	13
1	\checkmark	-	\checkmark	-	-	-		-	-		-	-	-
2	-	\checkmark	-	-	-	\checkmark	\checkmark	-	\checkmark	-	-	-	-
3		\checkmark	-	-	\checkmark	-	-	\checkmark	-	-	-	-	-
4	-	-	-	\checkmark	\checkmark	\checkmark	-	-	-	\checkmark	-	-	-
5	-	-	\checkmark	\checkmark	-	-	-	\checkmark	\checkmark	-	-	-	-
6	-	-	-	-	-	-	-	-	-	-	\checkmark	-	\checkmark
7	-	-	-	-	-	-	-	-	-	-	\checkmark	\checkmark	-
8	-	-	-	-	-	-	-	-	-	-	-	\checkmark	\checkmark

- also (fixed effects) collinearity between added reviewer and abstract effects,
- * reviewers 6-8 and abstracts 11-13 are separated from rest of design \Rightarrow abstracts cannot be compared to other abstracts without including effects of reviewers,
- * this type of collinearity may be less obvious visually, but can be detected in a fixed effects model.

Take-away message: it is probably useful to try a fixed effects model, in order to detect such potential collinearities so as to be aware of their impact on results.

RESULTS I (ORIGINAL DATA)

Good news: no collinearity (despite 2 reviewers with only 1 review).

Comparison of results from 3 approaches:

(1) simple means, (2) random effects model, (2m) mixed model (fixed reviewer effects).

Summary table for absolute rank differences between methods: mean (sd) below diagonal \ interquartile range (full range) above diagonal

Method	simple	random	mixed
simple	-	4.5-21 (0-59)	6-25.5(0-65)
random	13.9(11.3)	-	1-4 (0-22)
mixed	16.7 (12.8)	$3.2\;(3.3)$	-

Disagreements in classifications between methods, when splitting the abstracts 51:68: simple vs random: 7 + 7; simple vs mixed: 9 + 9; random vs mixed: 2 + 2.

Findings:

- $\circ\,$ differences substantial between simple and model-based, but relatively minor between the 2 models,
- inspection of the data reveals that simple and model-based ranks differ most when the 2 reviewers are both extreme in the same direction (both low, or both high).

EVALUATION OF METHODS (SIMULATION)

Aim: explore performance of methods (simple, random effects model, mixed model),

- \circ random variation across simulations: errors only, or errors + reviewer effects,
- \circ error variance: high (~ data) or low (about one-tenth),
- measures: mean (sd) of absolute rank differences, mean (sd) of misclassification 0 counts (for 51:68 split).

from	rom 100 simulations for each setting:							
	Setting	Measure	simple	random	mixed			
	error only	ranking diff.	21.7 (1.5)	18.8(1.6)	19.0(1.6)			
	$\sigma^2=223$	misclassif.	29.5 (3.7)	25.5(4.1)	26.0(4.1)			
	error only	ranking diff.	$16.7 \ (0.6)$	7.1 (0.6)	7.1 (0.6)			
	$\sigma^2=22$	misclassif.	20.0(2.4)	7.1~(2.3)	7.5~(2.4)			
	error + rev.	ranking diff.	21.1 (1.8)	18.4(1.5)	18.8(1.6)			
	$\sigma^2=223$	misclassif.	28.8(4.5)	25.1 (4.0)	26.1 (4.1)			
	error + rev.	ranking diff.	16.4(2.1)	7.2~(0.5)	7.2~(0.5)			
	$\sigma^2=22$	misclassif.	21.5(5.0)	$8.1 \ (2.5)$	8.2(2.4)			

Results from 100 simulations for each setting

- model-based methods perform clearly better with low error variance, but only slightly better with actual error variance,
- very minimal differences between random and mixed models, and random settings.

RESULTS II (AUGMENTED DATA)

Augmented data: additional review obtained for abstracts considered to be close to relevant cut-off and with clear disagreement in its two reviewer scores $\Rightarrow 253$ reviews (still 27 reviewers and 119 abstracts).

Impact of augmentation for the 15 abstracts involved: overall minor (average rank differences within (-8.3, 5)), and 5 changes in classification.¹

Summary table for absolute rank differences between methods: mean (sd) below diagonal \ interquartile range (full range) above diagonal

Method	simple	random	mixed		
simple	-	5-20(0-45)	6-25 (0-51)		
random	13.9(10.6)	-	1-4 (0-22)		
mixed	16.7(12.2)	$3.2\;(3.3)$	-		

Disagreements in classifications between methods, when splitting the abstracts 49:70: simple vs random: 7 + 7; simple vs mixed: 9 + 9; random vs mixed: 2 + 2 (same numbers as before, but not quite the same abstracts).

 $\circ~$ very similar results to those for the original data.

¹ Actual classification split for the augmented data was 49:70 instead of 51:68, so some differences may also be due the small change in proportions.

Some cautious conclusions:

- the model-based approach(es) seemed to work reasonably well, and only a simple scale change was required to meet model assumptions,
- some clear differences in rankings and classifications between model-based approach(es) and simple means, but simulation study showed the error variance to be too large to demonstrate one as clearly "more correct" for the data at hand,
- $\circ~$ augmentation of data with 15 additional reviews did not have much of an impact.

Additional methodological considerations:

- Bayesian modelling/estimation possible as well, but seems to agree well with (RE)ML estimation, and does not help with diagnosis of problems (results not shown here),
- only 2 reviews per abstract limits the options for more complex models:
 - $\ast\,$ congeneric measurement model (Skrondal & Rabe-Hesketh, 2004) is not identifiable,
 - * attempts to incorporate unequal variances across reviewers were not successful (convergence problems).

The \$1000 question is (of course):

What is the best approach to rank the next (and much larger) pool of abstracts?